

Some important inequalities

Probability and distributions

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Some important inequalities

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Some important
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We present some important (famous?) inequalities involving expectations

Some important inequalities (cont.)

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Theorem

Let X be a random variable and let m be a positive integer

Suppose $E(X^m)$ exists

If k is a positive integer such that $k \leq m$, then $E(X^k)$ exists

Proof: We prove it for the continuous case

Let $f(x)$ be the PDF of X , then

$$\begin{aligned}\int_{-\infty}^{\infty} |x|^k f(x) dx &= \int_{|x| \leq 1} |x|^k f(x) dx + \int_{|x| > 1} |x|^k f(x) dx \\ &\leq \int_{|x| \leq 1} f(x) dx + \int_{|x| > 1} |x|^m f(x) dx \\ &\leq \int_{-\infty}^{\infty} f(x) dx + \int_{-\infty}^{\infty} |x|^m f(x) dx \\ &\leq 1 + E(|X|^m) < \infty\end{aligned}\tag{1}$$



Some important inequalities (cont.)

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Theorem

Markov's inequality

Let $u(X)$ be a non-negative function of the random variable X

If $E[u(X)]$ exists, then for every positive constant c ,

$$P[u(X) \geq c] \leq \frac{E[u(X)]}{c}$$

Proof

We prove it for the continuous case

Let $A = \{x : u(x) \leq c\}$ and let $f(x)$ denote the PDF of X , then

$$E[u(X)] = \int_{-\infty}^{\infty} u(x)f(x) = \int_A u(x)f(x)dx + \int_{A^c} u(x)f(x)dx$$

Each of the integrals in the very RHS is non-negative

- The LHS is greater than or equal to either of them
- Specifically,

$$E[u(X)] \geq \int_A u(x)f(x)dx$$

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$$E[u(X)] \geq \int_A u(x)f(x)dx$$

If $x \in A$, then $u(x) \geq c$

Accordingly, the RHS does not increase if we replace $u(c)$ by c

$$E[u(X)] \geq c \int_A f(x)dx$$

Since,

$$\int_A f(x)dx = P(X \in A) = P[u(X) \geq c],$$

it follows that

$$E[u(X)] \geq cP[u(X) \geq c]$$



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The previous theorem generalises Chebyshev's inequality

Theorem

Chebyshev's inequality

Let the random variable X have a distribution of probability

Assume a finite variance $\text{Var}(X)$ ¹

Then, for every $k > 0$

$$P[|X - \mu| \geq k\sqrt{\text{Var}(X)}] \leq 1/k^2 \quad (2)$$

or, equivalently

$$P[|X - \mu| < k\sqrt{\text{Var}(X)}] \geq 1 - 1/k^2$$

¹This implies that the mean $\mu = E(X)$ exists (earlier theorem)

Some important inequalities (cont.)

Proof

In Markov's inequality, we can take $u(X) = (X - \mu)^2$ and $c = k^2 \text{Var}(X)$

$$P[(X - \mu)^2 \geq k^2 \text{Var}(X)] \leq \frac{E[(X - \mu)^2]}{k^2 \text{Var}(X)}$$

Since the numerator on the RHS is $\text{Var}(X)$, the inequality becomes

$$P[|X - \mu| \geq k\sqrt{\text{Var}(X)}] \leq 1/k^2$$

This is the desired result for a non-negative number k greater than one



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Example

Let X have the PDF

$$f(x) = \begin{cases} \frac{1}{2\sqrt{3}}, & x \in (-\sqrt{3}, \sqrt{3}) \\ 0, & \text{elsewhere} \end{cases}$$

Let $\mu = 0$ and $\sigma^2 = 1$

If $k = 3/2$, then we have the exact probability

$$P(|X - \mu| \geq k\sigma) = P\left(|X| \geq \frac{3}{2}\right) = 1 - \int_{-3/2}^{3/2} \frac{1}{2\sqrt{3}} dx = 1 - \sqrt{3}/2$$

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By Chebyshev's inequality, this probability is bounded from above

$$1/k^2 = 4/9$$

As $1 - \sqrt{3}/2 \approx 0.314$, the exact probability is smaller than the upper bound

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If $k = 2$, we have the exact probability

$$P(|X - \mu| \geq 2\sigma) = P(|X| \geq 2) = 0,$$

Again. much smaller than the upper bound from Chebyshev's inequality

$$1/k^2 = 1/4$$



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Example

Let the RV X of the discrete type have probabilities

- $\frac{1}{8}, \frac{6}{8}, \frac{1}{8}$, at the points $x = -1, 0, +1$, respectively

Let $\mu = 0$ and $\sigma^2 = 1/4$

If $k = 2$, then $1/k^2 = 1/4$ and $P(|X - \mu| \geq k\sigma) = P(|X| > 1) = 1/4$

The probability $P(|X - \mu| \geq k\sigma)$ reaches the upper bound $1/k^2 = 1/4$



- The inequality can not be improved
- Further assumptions about the distribution of X are needed

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Definition

A function ϕ defined on an interval (a, b) such that $-\infty \leq a < b \leq \infty$

Function ϕ is said to be **convex** if for all x, y in (a, b) and for all $\gamma \in (0, 1)$

$$\phi[\gamma x + (1 - \gamma)y] \leq \gamma\phi(x) + (1 - \gamma)\phi(y)$$

We say ϕ is **strictly convex** if the inequality is strict

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Theorem

If ϕ is differentiable on (a, b) , then

- *ϕ is convex iff $\phi'(x) \leq \phi'(y)$, for all $a < x < y < b$*
- *ϕ is strictly convex iff $\phi'(x) < \phi'(y)$, for all $a < x < y < b$*

If ϕ is twice-differentiable on (a, b) , then

- *ϕ is convex iff $\phi''(x) \geq 0$, for all $a < x < y < b$*
- *ϕ is strictly convex iff $\phi''(x) > 0$, for all $a < x < y < b$*

A very useful probability inequality follows from convexity

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Theorem

Jenses's inequality

Let ϕ be convex on an open interval I and X be a RV with support in I

Assume that X has a finite expectation

Then,

$$\phi[E(X)] \leq E[\phi(X)] \quad (3)$$

If ϕ is strictly convex, the inequality is strict, unless X is a constant RV

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Proof

We assume that ϕ has a second derivative (only convexity is needed)

Expand ϕ into a order 2 Taylor series about $\mu = E[X]$

$$\phi(x) = \phi(\mu) + \phi'(\mu)(x - \mu) + \frac{\phi''(\zeta)(x - \mu)^2}{2}$$

ζ is between x and μ

On the RHS, the last term on the RHS is nonnegative

$$\phi(x) \geq \phi(\mu) + \phi'(\mu)(x - \mu)$$

Taking expectations of both sides yields

Provided X is not a constant, the inequality is strict if $\phi''(x) > 0, \forall x \in (a, b)$



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Example

Let X be a non-degenerate RV with mean μ and a finite second moment

↪ Then, $\mu^2 < E(X^2)$

Obtained by Jensen's inequality using a strictly convex function $\phi(t) = t^2$



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Example

Harmonic and geometric means

Let $\{a_1, \dots, a_n\}$ be a set of positive numbers

Set a distribution for RV X by placing mass $1/n$ on each of the numbers a_i

Then, the mean of X is the **arithmetic mean (AM)**

$$E(X) = n^{-1} \sum_{i=1}^n a_i$$

Since $-\log(x)$ is a convex function, by Jensen's inequality, we have

$$-\log\left(\frac{1}{n} \sum_{i=1}^n a_i\right) \leq E[-\log(X)] = -n^{-1} \sum_{i=1}^n \log(a_i) = -\log(a_1 a_2 \cdots a_n)^{1/n}$$

or, equivalently

$$\left(\frac{1}{n} \sum_{i=1}^n a_i\right) \geq (a_1 a_2 \cdots a_n)^{1/n}$$

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Hence,

$$(a_1 a_2 \cdots a_n)^{1/n} \leq \frac{1}{n} \sum_{i=1}^n a_i \quad (4)$$

The quantity on the LHS is the **geometric mean (GM)**

- For any finite set of positive numbers, $\text{GM} \leq \text{AM}$

Some important inequalities (cont.)

By replacing a_i by its reciprocal $1/a_i$ (also positive),

$$\frac{1}{n} \sum_{i=1}^n \frac{1}{a_i} \geq \left(\frac{1}{a_1} \frac{1}{a_2} \cdots \frac{1}{a_n} \right)^{1/n}$$

or equivalently

$$\frac{1}{1/n \sum_{i=1}^n 1/a_i} \leq (a_1 a_2 \cdots a_n)^{1/n}$$

The quantity on the LHS is the **harmonic mean (HM)**



Remark

Putting things together

- For any finite set of positive numbers, $\text{HM} \leq \text{GM} \leq \text{AM}$