

The binomial and the Poisson distribution

Useful distributions

Francesco Corona

Department of Computer Science
Federal University of Ceará, Fortaleza

The binomial
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Poisson
distribution

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The binomial
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The binomial distribution

Useful distributions

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A **Bernoulli experiment** is a random experiment

Its outcome can be classified in one of two mutually exclusive ways

- Success or failure, defective and non-defective
- Life or death, female or male
- ...

The binomial distribution (cont.)

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A sequence of **Bernoulli trials**

- ↪ A Bernoulli experiment performed several independent times
- ↪ The probability of **success** remains the same

Let p indicate the probability of **success** on each trial

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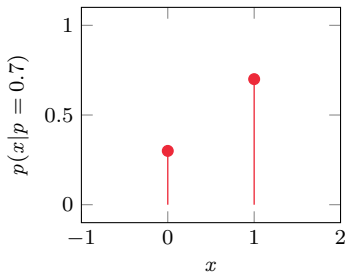
The Poisson
distribution

Let X be a random variable associated with a Bernoulli trial

- $X(\text{success}) = 1$
- $X(\text{failure}) = 0$

The two outcomes, **success/failure**, are indicated by 0/1

The PMF of the random variable X



$$p(x) = p^x(1-p)^{1-x} \quad (1)$$

for $x = 0, 1$

X is said to have a **Bernoulli distribution**, with parameter p

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The expected value of X

$$\begin{aligned}\mu = E(X) &= \sum_{x \in \{0,1\}} xp(x) = \sum_{x \in \{0,1\}} xp^x(1-p)^{(1-x)} \\ &= (0)(1-p) + (1)(p) \\ &= p\end{aligned}$$

The variance of X

$$\begin{aligned}\sigma^2 = \text{var}(X) &= p^2(1-p) + (1-p)^2p \\ &= p(1-p)\end{aligned}$$

The standard deviation of X

$$\sigma = [p(1-p)]^{1/2}$$

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A sequence of n Bernoulli trials

Suppose that n independent Bernoulli trials are performed

- Each with the same success probability

An observed sequence of n trials is a n -tuple of 0s and 1s

$0, 0, 1, 0, 1, 1, 1, 0, \dots$

n times

$1, 0, 1, 0, 0, 0, 1, 0, \dots$

n times

$1, 1, 0, 1, 0, 1, 0, 0, \dots$

n times

There are a total of 2^n possible such sequences¹

¹Among them, those without repetitions (permutations) are $n!/(n-n)! = n!$.

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We are interested in the distribution of the number of successes in n trials

- The order of their occurrence is not of our interest

The number of sequences with a single 1 (success) out of n trials

$$\binom{n}{1} = \frac{n!}{1!(n-1)!} = n$$

The number of sequences with two 1s (successes) out of n trials

$$\binom{n}{2} = \frac{n!}{2!(n-2)!} = n(n-1)/2$$

The number of sequences with x 1s (successes) out of n trials

$$\binom{n}{x} = \frac{n!}{x!(n-x)!}$$

If we observe x successes ($x = 0, 1, 2, \dots, n$), then also $(n-x)$ failures occur

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Consider a sequence of n Bernoulli trials

Let the number of observed successes in n trials be the random variable X

- The possible values of X are $x = 0, 1, 2, \dots, n$

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Let X_i denote the Bernoulli random variable associated with the i -th trial

- The probabilities of success and failure on each trial are p and $1 - p$

$$p^x(1-p)^{1-x}, \text{ for } x = \{0, 1\}$$

The probability of any of the possible ways to get x successes in n trials

↪ The Bernoulli random variables are assumed independent

$$\sum_{x=0}^n p^x(1-p)^{1-x} = p^x(1-p)^{n-x}$$

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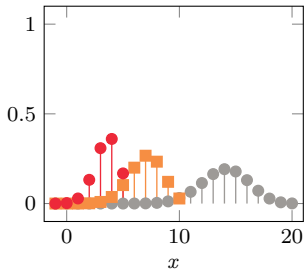
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$$p(x|p = 0.7, n = 5|10|20)$$



$$p(x) = \begin{cases} \binom{n}{x} \frac{p^x}{(1-p)^{x-n}}, & x = 0, 1, \dots, n \\ 0, & \text{elsewhere} \end{cases} \quad (2)$$

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Remark

If n is a positive integer

$$(a + b)^n = \sum_{x=0}^n \binom{n}{x} b^x a^{n-x}$$

Thus,

$$\begin{aligned} \sum_x p(x) &= \sum_{x=0}^n \binom{n}{x} p^x (1-p)^{n-x} \\ &= [(1-p) + p]^n \\ &= 1 \end{aligned}$$

It is clear that $p(x) \geq 0$

$p(x)$ satisfies the conditions of being a PMF of a RV X of the discrete type

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A RV X with the PMF of the form of $p(x)$ has a **binomial distribution**

- Any such $p(x)$ is called a **binomial PMF**, with parameters n and p

A binomial PMF is denoted by the symbol $b(n, p)$

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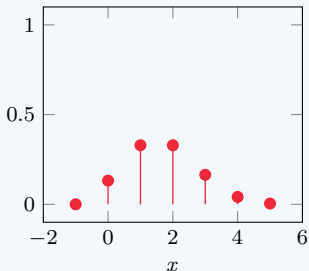
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Example

If we say that $X \sim b(5, 1/3)$, we mean that X has the binomial PMF

$$p(x|p = 1/3, n = 5)$$



$$p(x) = \begin{cases} \binom{5}{x} \frac{(1/3)^x}{(2/3)^{x-5}}, & x = 0, 1, \dots, 5 \\ 0, & \text{elsewhere} \end{cases} \quad (3)$$



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The MGF of a binomial distribution

$$\begin{aligned}M(t) &= \sum_x e^{(tx)} p(x) = \sum_{x=0}^n e^{(tx)} \binom{n}{x} p^x (1-p)^{n-x} \\ &= \sum_{x=0}^n \binom{n}{x} [pe^{(t)}]^x (1-p)^{n-x} \\ &= [(1-p) + e^t]^n\end{aligned}$$

for all real values of t

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The mean μ and the variance σ^2 of X may be computed from $M(t)$

$$M'(t) = n[(1-p) + pe^{(t)}]^{n-1} pe^{(t)}$$

$$M''(t) = n[(1-p) + pe^{(t)}]^{n-1} pe^{(t)} + \\ n(n-1)[(1-p) + pe^{(t)}]^{n-2} [pe^{(t)}]^2$$

Thus,

$$\mu = M'(0) = np$$

$$\sigma^2 = M''(0) = np + n(n-1)p^2 - (np)^2 = np(1-p)$$

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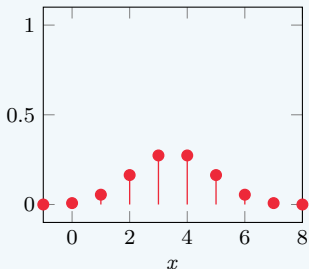
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Example

Let X be the number of successes in $n = 7$ independent fair experiments

The PMF of X

$$p(x|p = 0.5, n = 7)$$



$$p(x) = \begin{cases} \binom{7}{x} \frac{(1/2)^x}{(1/2)^{x-7}}, & x = 0, 1, \dots, 7 \\ 0, & \text{elsewhere} \end{cases} \quad (4)$$

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The MGF of X

$$M(t) = (1/2 + 1/2e^t)^7$$

The mean of X

$$\rightsquigarrow \mu = np = 7/2$$

The variance of X

$$\rightsquigarrow \sigma^2 = np(1 - p) = 7/4$$

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Furthermore,

$$P(0 \leq X \leq 1) = \sum_{x=0}^1 p(x) = \frac{1}{128} + \frac{7}{128} = 8/128$$

$$P(X = 5) = p(5) = \frac{7!}{5!2!} (1/2)^5 (1/2)^2 = 21/128$$



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The commands to obtain binomial probabilities for $X \sim b(n, p)$

- `dbinom(k,n,p)`, $P(X = k)$
- `pbinom(k,n,p)`, $P(X \leq k)$

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Example

Let X be a random variable with the MGF

$$M(t) = (2/3 + 1/3e^t)^5$$

Then, X has a binomial distribution with $n = 5$ and $p = 1/3$

The binomial distribution (cont.)

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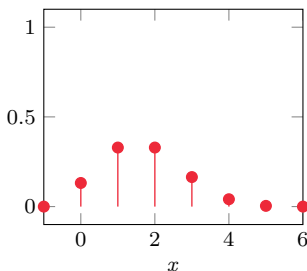
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The PMF of X

$$p(x|p = 1/3, n = 5)$$



$$p(x) = \begin{cases} \binom{5}{x} \left(\frac{1}{3}\right)^x \left(\frac{2}{3}\right)^{5-x}, & x = 0, 1, \dots, 5 \\ 0, & \text{elsewhere} \end{cases}$$

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The mean of X

$$\mu = np = 5/3$$

The variance of X

$$\sigma^2 = np(1 - p) = 10/9$$



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Example

Let the random variable Y be equal to the number of successes out of n independent random experiments with probability of success equal p

The ratio Y/n is called the **relative frequency of success**

Recalling Chebyshev's inequality, for all $\varepsilon > 0$, we have

$$P\left(\left|\frac{Y}{n} - p\right| \geq \varepsilon\right) \leq \frac{\text{Var}(Y/n)}{\varepsilon^2} = \frac{p(1-p)}{n\varepsilon^2}$$

For any fixed ε , the RHS goes to zero with n

$$\lim_{n \uparrow \infty} P\left(\left|\frac{Y}{n} - p\right| \geq \varepsilon\right) = 0$$

The Binomial and related distributions (cont.)

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Thus,

$$\lim_{n \uparrow \infty} P\left(\left|\frac{Y}{n} - p\right| < \varepsilon\right) = 1$$

The relative frequency of success is close to the probability of success p

- For sufficiently large n
- (And any $\varepsilon > 0$)

The result is a form of the **Weak Law of Large Numbers (WLLN)**



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Theorem

Let X_1, X_2, \dots, X_m be independent random variables

Let each X_i have the binomial distribution $b(n_i, p)$, for $i = 1, 2, \dots, m$

Let $Y = \sum_{i=1}^m X_i$

Then, Y has a binomial distribution $b(\sum_{i=1}^m n_i, p)$

Proof

The MGF of X_i

$$M_X(t) = [(1-p) + pe^{(t)}]^{n_i}$$

By independence,

$$M_Y(t) = \prod_{i=1}^m [(1-p) + pe^{(t)}]^{n_i} = [(1-p) + pe^{(t)}]^{\sum_{i=1}^m n_i}$$

Hence, Y has the binomial distribution $b(\sum_{i=1}^m n_i, p)$



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Recall² the series

$$1 + m + \frac{m^2}{2!} + \frac{m^3}{3!} + \cdots = \sum_{x=0}^{\infty} \frac{m^x}{x!}$$

The series converges absolutely, for *all* values of m , to $\exp(m)$

Consider the function $p(x|m)$, for some $m > 0$

$$p(x|m) = \begin{cases} \frac{m^x}{x!} e^{-m}, & x = 0, 1, 2, \dots \\ 0, & \text{elsewhere} \end{cases} \quad (5)$$

²Walter Rudin. *Real and complex analysis*, page 1.

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Since $m > 0$, then $p(x) \geq 0$

Moreover,

$$\begin{aligned}\sum_x p(x) &= \sum_{x=0}^{\infty} \frac{m^x e^{(-m)}}{x!} = e^{(-m)} \sum_{x=0}^{\infty} \frac{m^x}{x!} \\ &= e^{(-m)} e^{(m)} = 1\end{aligned}$$

$p(x)$ satisfies the condition of being a PMF of a discrete type of RV

The Poisson distribution (cont.)

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A RV X with the PMF of the form $p(x)$ has a **Poisson distribution**

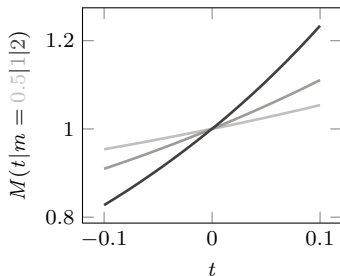
- Any such $p(x)$ is called a **Poisson PMF** with parameter m

The Poisson distribution (cont.)

The MGF of a Poisson distribution

$$\begin{aligned}M(t) &= \sum_x e^{tx} p(x) = \sum_{x=0}^{\infty} e^{tx} \frac{m^x e^{-m}}{x!} \\&= e^{-m} \sum_{x=0}^{\infty} \frac{(me^t)^x}{x!} \\&= e^{-m} e^{me^t} = e^{m(e^t-1)}\end{aligned}$$

for all real values of t



The Poisson distribution (cont.)

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We have,

$$M'(t) = e^{m(e^t-1)}(me^t)$$

$$M''(t) = e^{m(e^t-1)}(me^t) + e^{m(e^t-1)}(me^t)^2$$

Thus,

$$\rightsquigarrow \mu = M'(0) = m$$

$$\rightsquigarrow \sigma^2 = M''(0) - \mu^2 = m + m^2 - m^2 = m$$

The Poisson distribution has $\mu = \sigma^2 = m > 0$

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An alternative form of the Poisson PMF

$$p(x) = \begin{cases} \frac{\mu^x e^{-\mu}}{x!}, & x = 0, 1, 2, \dots \\ 0, & \text{elsewhere} \end{cases}$$

Thus, the parameter μ in a Poisson PMF is the mean μ

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Let X be a RV with the Poisson distribution with parameter $m = \mu$

Then,

- `dpois(k,m)`, $P(X = k)$
- `ppois(k,m)`, $P(X \leq k)$

The Poisson distribution (cont.)

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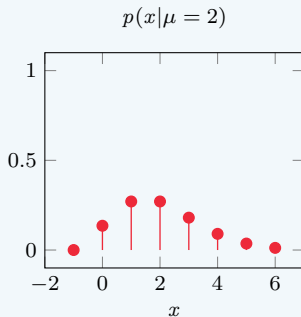
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Example

Suppose that X has a Poisson distribution with $\mu = 2$, then the PMF of X



$$p(x) = \begin{cases} \frac{2^x e^{-2}}{x!}, & x = 0, 1, 2, \dots \\ 0, & \text{elsewhere} \end{cases}$$

The Poisson distribution (cont.)

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The variance of this distribution is $\sigma^2 = \mu = 2$

We can compute $P(1 \leq X)$ or use the table of distributions

$$\begin{aligned}P(1 \leq X) &= 1 - P(X = 0) \\ &= 1 - p(0) = 1 - e^{-2} \\ &= 0.865\end{aligned}$$



The Poisson distribution (cont.)

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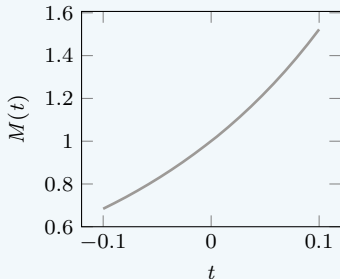
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Example

Let the random variable X have the MGF



$$M(t) = e^{4(e^t - 1)}$$

Then, X has a Poisson distribution

Accordingly,

$$P(X = 3) = \frac{4^3 e^{-4}}{3!} = \frac{32}{3} e^{-4}$$

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By the table of distributions,

$$P(X = 3) = P(X \leq 3) - P(X \leq 2) = 0.433 - 0.238 = 0.195$$



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The Poisson distribution satisfies an important additive property

Theorem

Suppose X_1, \dots, X_n are independent random variables

Suppose that each X_i has a Poisson distribution with parameter m_i

Then, $Y = \sum_{i=1}^n X_i$ has a Poisson distribution with parameter $\sum_{i=1}^n m_i$

Proof

We obtain the result by determining the MGF of Y

$$\begin{aligned}M_Y(t) &= E[\exp(tY)] = \prod_{i=1}^n \exp\{m_i[\exp(t) - 1]\} \\ &= \exp\left\{\sum_{i=1}^n m_i[\exp(t) - 1]\right\}\end{aligned}$$

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By the uniqueness of MGFs, Y has a Poisson distribution with parameter

$$\rightsquigarrow \mu = \sum_{i=1}^n m_i$$

